

## Review of chapters 8 and 9

We will go through these examples in lecture, Thursday November 29th.

1. Let  $X_1, \dots, X_n$  be a random sample from  $N(0, \sigma^2)$ . Show that  $\sum_{i=1}^n X_i^2/n$  is an unbiased estimator of  $\sigma^2$ .
2. The following problems from the book: 8.2.9, 8.2.10, 8.4.3, 8.6.15, 9.8.3

### Other good problems

Notation:  $\text{No}(\mu, \sigma^2)$  is the normal distribution.

**Problem 2:** The  $n = 9$  independent random variables  $\{X_i\}_{i \leq n}$  are a simple random sample from the Normal distribution  $X_i \sim \text{No}(\mu, \sigma^2)$ , with variance  $\sigma^2$  and mean  $\mu$ .

- a) (7) With  $\sigma^2$  unknown, describe the Rejection Region  $\mathcal{R}$  for a two-sided test of size  $\alpha = 0.05$  of the hypothesis

$$H_0 : \mu = 17.5 \quad \text{vs.} \quad H_1 : \mu \neq 17.5$$

Reject  $H_0$  if:

- b) (7) Same question, if variance of  $\{X_i\}$  is known to be  $\sigma^2 = 16$ :

Reject  $H_0$  if:

- c) (6) What is the *power*  $\pi(\mu)$  of the test in part b) above, for  $\mu = 20$ ?

$$\pi(20) = \underline{\hspace{2cm}}$$

**Problem 3:** Once again the  $n = 9$  independent random variables  $\{X_i\}_{i \leq n}$  are a simple random sample from the Normal distribution  $X_i \sim \text{No}(\mu, \sigma^2)$ , with variance  $\sigma^2$  and mean  $\mu$ , but now we have some data. The values of a few statistics from this random sample are:

$$S(\mathbf{x}) = \sum_{i \leq n} X_i = 180 \quad T(\mathbf{x}) = \min_{i \leq n} X_i = 1$$

$$V(\mathbf{x}) = \sum_{i \leq n} (X_i - \bar{X})^2 = 72 \quad W(\mathbf{x}) = \text{Median}(\{X_i\}) = 24$$

a) (7) With  $\sigma^2$  unknown, give the  $P$ -value for a two-sided test of the hypothesis

$$H_0 : \mu = 17.5 \quad \text{vs.} \quad H_1 : \mu \neq 17.5$$

$P =$

b) (7) Same question, if variance of  $\{X_i\}$  is known to be  $\sigma^2 = 16$ :

$P =$

c) (6) Are these answers consistent with your answers to parts a) and b) of Problem 2? **Y N** Explain.

d) With  $\sigma^2$ , give the 95% symmetric two-sided confidence interval for  $\mu$

For the next problem note that if  $p$  is small we expect  $X$  to be large and vice versa, so for a) you should use a test that rejects  $H_0$  if  $X \geq c$  for some  $c$ . Also, the hypotheses in a) and b) are in fact

$$a) \quad H_0 : p \geq 1/2 \quad \text{vs.} \quad H_0 : p < 1/2$$

and

$$b) \quad H_0 : p \leq 1/2 \quad \text{vs.} \quad H_0 : p > 1/2$$

**Problem 8:** A series of Statistics students answer (independently, of course) the question “Was your statistics final too long?” We record the number  $X$  of consecutive “No, it was fine” answers before the first “Well, I thought it was a little too long.” Of course  $X$  has the Geometric Distribution, with probability mass function

$$f(x | p) = P[X = x | p] = p q^x, \quad x = 0, 1, 2, \dots$$

We wish to test the hypothesis  $H_0 : p = 1/2$  about the probability  $p$  of a “Too long” answer. If  $H_0$  were true, then “Yes” and “No” answers would be like Heads and Tails of a fair coin. Give all answers **exactly** and in as **simple** a form as you can. For parts a) and b) we observe  $X = 3$ :

- a) (5) Find the  $P$ -value for  $H_0 : p = 1/2$  vs.  $H_1 : p < 1/2$ , if  $X = 3$   
[First decide what outcomes are more extreme for this  $H_1$ ]:

$$P = \underline{\hspace{2cm}}$$

- b) (5) Now for  $H_0 : p = 1/2$  vs. the other one-sided alternative  $H_1 : p > 1/2$ , if  $X = 3$ :

$$P = \underline{\hspace{2cm}}$$

**Problem 8 (cont):**

Now let's be more optimistic and suppose we observe  $X = 10$ :

- c) (2) For  $H_0 : p = 1/2$  vs.  $H_1 : p < 1/2$ , if we observe  $X = 10$ :

$$P = \underline{\hspace{2cm}}$$

- d) (8) With a uniform prior distribution with density  $\xi(p) = \mathbf{1}_{\{0 < p < 1\}}$ , what is the posterior expectation of  $p$  if we observe  $X = 10$ ?

$$E_{\xi}(p | x = 10) = \underline{\hspace{2cm}}$$